RESEARCH ARTICLE

WILEY

Promoting academic satisfaction and performance: Building academic resilience through coping strategies

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Funding information

Fundació Caixa Castelló-Bancaixa, Grant/ Award Number: E 2013-31; Consellería de Educación Generalitat Valenciana, Grant/ Award Number: PROMETEO/2013/025

Abstract

The aim of this article is twofold: (a) to validate the psychometric properties of the Academic Resilience Scale (AR-S) in a Spanish university context; and (b) to test a model where different coping strategies are antecedents of AR, and where academic satisfaction and performance are its consequences. The studies were conducted with 185 (study 1) and 780 (study 2) students. In study 1, the construct validity, reliability, and convergent and divergent validity of the AR-S were analyzed. The hypothesized model (study 2) was tested using structural equation modeling. The results confirmed the good psychometric properties of the AR-S in the Spanish context, providing evidence of its validity. Furthermore, the distinctive role of different coping strategies in resilience was confirmed, as well as the relationship between resilience and academic satisfaction. No direct relationship between resilience and performance was found, highlighting the full mediation of satisfaction.

KEYWORDS

academic performance, academic resilience, academic satisfaction, coping, scale validation

1 | INTRODUCTION

In light of the increasing need for healthy and competent workers who can cope with the current dynamic and changing work context, it is crucial to educate and motivate university students to be valuable and contributors to the economy and the nation in the long run (Riolli, 2012). Research has shown that psychological distress is

significantly higher in university students than in the general population (Stallman, 2010), and it has also been established that stress increases as students progress in their studies (Putwain, 2007). Indeed, university students face many challenges in the pursuit of their educational goals, including exams, demands of their academic course work, and managing intrapersonal and interpersonal changes (Roddenberry & Renk, 2010; Ross, Niebling, & Heckett, 1999). Although these challenges may be considered microstressors, due to their accumulated effects over time, they are important for health (Zautra, 2003) and can have a variety of negative psychological and behavioral effects that can reduce students' performance. For instance, academic stress has been associated with disorders such as depression (Dao, Lee, & Chang, 2007), school failure (McDonald, 2001), study burnout (Salanova, Schaufeli, Martínez, & Bresó, 2010), and immune system disturbances (Vedhara & Nott, 1996).

However, some students have the capacity to successfully withstand the consequences of negative academic experiences. They are easily encouraged following minor setbacks and generally view negative challenging events as surmountable (e.g., Struthers & Perry, 1996). Nevertheless, little is known about the relative strength of various factors that promote positive adaptation in the face of academic stress (Leary & DeRosier, 2012). The concept of resilience might help to explain why some individuals who experience high levels of stress are able to withstand it and may even thrive on it, enabling them to manage future challenges effectively (Grant & Kinman, 2012). Thus, to account for the discrepancy in students' responses to negative events, the present study focused on academic resilience (AR) as a factor that promotes positive adaptation and outcomes during the academic stage.

From a stress-coping perspective, numerous studies have highlighted the importance of coping in dealing with adversity (e.g., Clauss-Ehlers, 2008; Leipold & Greve, 2009), and individuals use a range of coping strategies for this purpose. Therefore, in stressful academic-related situations, students may have different ways to face them; that is, they possess different coping strategies that will have different effects on well-being and performance. In this study, we investigate the predictive role of distinct coping strategies on AR. Moreover, we examine whether the relationship between coping strategies and AR stimulates positive academic outcomes, assessed by two widely recognized indicators, namely, academic satisfaction and performance. In sum, this study aims to further our understanding about how coping strategies promote resilience, leading to greater satisfaction and better performance.

This study extends previous research in this field in several ways. First, a validation of an AR scale in the Spanish context is proposed to have a reliable instrument to measure resilience and test for its antecedents and outcomes. Second, although earlier studies have examined coping strategies as predictors of resilience, they mainly combined scales into general problem- and emotion-focused categories of coping. Instead, as Carver, Scheier, and Weintraub (1989) suggest, we use the data to determine the composition of the higher-order factors because different samples exhibit different patterns of relations. Third, we include the Grade Point Average (GPA) provided by the University as an objective measure of academic performance to control for method bias, thus strengthening the validity of our results. Moreover, academic performance (GPA) is measured 4 or 5 months after the administration of the questionnaire with the psychosocial variables, thus allowing us to study the effect of psychosocial variables on performance over time. Finally, we use a sample of students pursuing different degrees in different faculties to reinforce the generalizability of the results.

1.1 | About the concept and measure of AR

It has been argued that focusing on salutogenic (or health-enhancing) rather than pathogenic (or disease-reducing) factors is more successful in promoting well-being in various life contexts (Seligman, 2003). Based on this framework, in this study we focus on resilience because it refers to the process of positive adaptation in the face of stress or adversity (Luthar, Cicchetti, & Becker, 2000). Applied to the academic context, resilience research focuses on the mechanisms and interactive processes that enable students to succeed despite unfavorable circumstances (Doll & Lyon, 1998). We refer to AR as the process of dealing with academic adversity and achieving positive outcomes in stressful situations. A substantial body of evidence reports that resilient students are those who

sustain high levels of achievement motivation and performance, despite the presence of stressful events and conditions (Leary & DeRosier, 2012). To date, the few studies that have dealt with AR have focused on ethnic-minority groups, learning difficulties, and extreme underachievers. However, AR is relevant to all students because at some point they all experience challenges during their academic life (Martin & Marsh, 2006). Consequently, we propose that the study of AR can help to further our understanding of the process through which students are able to successfully adapt to and "bounce back" from stressful situations, as well as providing insight into possible interventions (Steinhardt & Dolbier, 2008).

It is interesting to note that in recent years different instruments have been used to measure resilience in student populations; however, in most cases they are not specific to the academic context (e.g., Campbell-Sills, Cohan, & Stein, 2006; Dolbier, Jaggars, & Steinhardt, 2010). To our knowledge, the scale proposed by Martin and Marsh (2006) is the only one that: (a) focuses specifically on AR, and (b) has different content from previous measures, which have generally assessed protective factors or resources involving personal characteristics (e.g., Connor & Davidson, 2003), framing its items as statements referring to the process of bouncing back, adapting to stress, or thriving in the face of academic adversity. Thus, we aim to evaluate the psychometric characteristics of the Spanish version of the AR-S by Martin and Marsh (2006) in the first study in this paper (see study 1). After confirming that the AR-S is a reliable scale to measure AR in the Spanish context, it can be used as a measure of resilience to test antecedents and outcomes.

1.2 | Coping strategies as antecedents of AR

To understand people who positively face adversity, it is undoubtedly important to identify the factors that may promote resilience, such as coping strategies (Smith et al., 2008; Steinhardt & Dolbier, 2008). Coping has been defined as a person's efforts to remove, reduce, or manage threatening events or situations that are appraised as challenging or stressful (Lazarus & Folkman, 1984). Although resilience and coping are sometimes used interchangeably, they are in fact related, but conceptually distinct, constructs. The distinction between AR and coping strategies is based on the definition of the constructs: whereas coping refers to the set of cognitive and behavioral strategies used by an individual to manage the demands of stressful situations, resilience refers to adaptive outcomes in the face of adversity (Campbell-Sills et al., 2006).

Researchers have generally clustered coping responses into theoretically derived factors. One of the earlier nomenclatures (Folkman & Lazarus, 1980) distinguishes two major functions of coping: problem-focused coping and emotion-focused coping. Problem-focused coping involves addressing the problem causing the distress. Effective problem-focused coping probably contributes to positive psychological states by allowing people to experience some personal control and a sense of accomplishment. Emotion-focused coping is aimed at regulating distress and negative emotion, rather than trying to change the events themselves, through the use of strategies such as escape-avoidance or support-seeking. These strategies involve thoughts and/or actions that relieve or lessen the emotional impact of stress. Research has found that people who cope with stress by seeking social support or venting their feelings (emotion-focused) experience more negative outcomes than people who address the experienced stressor directly by working on solving their problems (problem-focused; Suldo, Shaunessy, Michalowski, & Shaffer, 2008). Previous studies on coping strategies and resilience also confirmed these results (Campbell-Sills et al., 2006).

Studies found that the problem-focused and emotion-focused distinction was a good starting point, but they also identified meaning-focused coping as a different type of coping where cognitive strategies are used to manage the meaning of a situation (Folkman & Moskowitz, 2004). Meaning-focused coping is basically aimed at regulating positive emotions. There is considerable evidence suggesting that positive emotions occur alongside negative emotions throughout intensely stressful periods and are of great significance for adaptation (Folkman, 2008). Therefore, positive emotions need to be included to learn more about how people generate and sustain them, and explore their adaptation role in relation to outcomes.

Drawing on these three different kinds of coping categories, we postulate distinct hypotheses about the relationship between coping strategies and resilience. First, we propose that students who show high levels of active coping and use problem-focused strategies are more resilient. Second, we suggest that students who use emotion-focused strategies, which are aimed at regulating distress and negative emotions, reduce their efforts in dealing with the problem and, thus, are less resilient. Finally, we propose that students who regulate positive emotions by attributing meaning to the situation broaden their thought-action repertoires and build resources, such as resilience, that enable them to cope and manage things effectively.

1.3 | Consequences of AR (satisfaction and performance)

In studying the consequences of AR, we focus on academic satisfaction and academic performance. In this study, academic satisfaction is considered students' cognitive evaluations of various aspects of their academic context (i.e., teachers, degree, and faculty). Academic satisfaction is considered the subjective intrinsic gratification students experience through their educational pursuits (Ryan, 2001), and its importance as an indicator of school adjustment (Baker, Dilly, Aupperlee, & Patil, 2003) and quality in higher education (Byrne & Flood, 2003) has been recognized. On the grounds that AR is useful in helping students to overcome the stress in the academic context, we propose that, when students show a high level of resilience, they will experience higher satisfaction with the aspects of their academic context. Indeed, students with higher levels of resilience, even experiencing adverse events in their academic context, are expected to positively adapt to and successfully bounce back from these events (Leary & DeRosier, 2012). Therefore, we postulate that this positive response enhances their academic satisfaction.

Previous research has suggested that AR leads to an increase in academic performance (Kwok, Hughes, & Wen, 2007) because it prepares individuals to rebound or bounce back from adversities, problems, and failures. Resilient people are more flexible in responding to changing demands, and they tend to use setbacks as "springboards" or opportunities for growth (Tugade & Fredrickson, 2004). However, we expect resilience to influence academic performance through satisfaction as well. Specifically, we propose that, when resilience is high, students perform better because they experience more academic satisfaction engendered by resilience. Previous evidence showed that attitudes are proximal antecedents and guidelines for behavior (Ajzen & Fishbein, 2005), and studies have shown that a cognitive attitude (i.e., satisfaction) is the primary source of student performance (Chambel & Curral, 2005; Rode et al., 2005). Hence, we argue that academic satisfaction is a partial mediator of the effects of AR on academic performance.

2 | STUDY 1

The first study assessed the psychometric properties of a Spanish adaptation of the Academic Resilience Scale (AR-S) developed by Martin and Marsh (2006). We conducted a number of tests to determine the validity and reliability of the measure. First, a confirmatory factor analysis was carried out, and construct validity was tested. Subsequently, the reliability of the AR-S was investigated, along with its convergent and divergent validity. Thus, we predict:

Hypothesis 1. The Spanish version of the AR-S will demonstrate acceptable psychometric properties (i.e., validity and reliability).

2.1 | Method

2.1.1 | Sample and procedures

The first study involved 185 students from a Spanish university. The sample size fits all the criteria suggested by Muthén and Muthén (2002) for determining sample size through a Monte Carlo study: (a) parameter and standard

error biases did not exceed 10% for any parameter in the model; (b) the standard error biases for parameters that are the specific focus of the power analysis (e.g., the factor covariance of resilience and self-efficacy) did not exceed 5%; and (c) coverage, that is, the confidence intervals, was between 0.91 and 0.98. In addition to these criteria, an appropriate sample size was determined because the power of the salient model parameters was above 0.80.

Participants (57.5% females) were stratified and belonged to the four faculties of the University, that is: Faculty of Humanities and Social Sciences (56.5%), School of Technology and Experimental Sciences (17.4%), Faculty of Health Sciences (14.7%), and Faculty of Law and Economics (11.4%). Ages ranged from 18 to 44 years (M = 22.8; SD = 5.06 years). Most of them (95.6%) were studying a bachelor's degree lasting 4 years, whereas the others were studying five-year degrees from the previous curriculum (4.4%). Regarding the course they were taking, 13.7% were in their first year, 50.5% in the second, 27.5% in the third, 7.1% in the fourth, and 1.1% in their fifth year. Eighty percent of them were not working at that moment.

The data were collected through a paper and pencil questionnaire that includes all the study variables. Students received a brief presentation of the study by the researchers during class time, and they were invited to individually fill out a questionnaire. On the cover page of the questionnaire, they read and signed an individual informed consent. The data collected by the questionnaire and the data on academic performance (GPA provided by the University) were matched through the National Identity document of each student.

2.1.2 | Measures

All the items on the measures used were scored on a 7-point scale ranging from 0 (*strongly disagree*) to 6 (*strongly agree*). On the questionnaire, the order of the items on the four measures was alternated.

The Academic Resilience Scale

The Spanish version of the AR-S was a translation of the Martin and Marsh (2006) scale. It is one-dimensional and includes 6 items, each framed in a positive direction. An example item is: "I think I'm good at dealing with schoolwork pressures". The AR score was obtained by computing the mean of the 6 items on the scale. The original α for this scale was 0.89 (Martin & Marsh, 2006). The translation was performed by two researchers who were fluent in Spanish and English, and then two other researchers were asked to compare the Spanish and English versions of the scale.

Other measures

A number of existing measures were used in this study to assess the convergent, divergent, and construct validity of the AR-S. For each of them, the score was obtained by computing the mean of the items that make up each scale.

Resilience Another measure used to assess resilience was the short 12-item version of the Psychological Capital Questionnaire (PCQ-12; see Luthans, Avey, Smith, & Li, 2008), adapted to the academic context. In this instrument, resilience was measured with a three-item scale. An example item is: "I usually take stressful things in my stride with regard to my studies". Cronbach's α was 0.60.

Self-efficacy It was measured with a five-item scale developed by Midgley and colleagues (2000) and used with a sample of Spanish students (Salanova, Bresó, & Schaufeli, 2005). An example item is: "I can do even the hardest work in this class if I try". Cronbach's α was 0.74.

Proving goal orientation Proving goal orientation was assessed with a four-item scale from the Goal Orientation scale developed by Vandewalle (1997). An example item is: "It's important for me to prove that I am better than others in class". Cronbach's α was 0.75.

2.2 Results

2.2.1 | Construct validity

With the purpose of confirming the one-factor structure of the AR-S and its relatedness to other constructs, we performed a confirmatory factor analysis, including a theoretically relevant concept, self-efficacy. From a theoretical point of view, self-efficacy has been proposed as an important supportive characteristic of AR. Bandura (2011) suggested that people who are high in self-efficacy view impediments as surmountable by developing required competencies and making a perseverant effort, and so they are able to face difficulties and remain resilient to adversity. Thus, it is proposed that students' beliefs about themselves and their academic capacities influence their behaviors and emotional reactions, allowing them to perform successfully in the face of obstacles or adverse experiences.

To shed more light on the relationship between these variables, structural equation modeling (SEM) was used to test the fit of the structural models. The analysis was performed using AMOS 21.0 (Arbuckle, 2005), and various fit indices were used to assess the model fit. Goodness of fit can be determined with the following indices (Hu & Bentler, 1999): the chi-square test (χ^2); Comparative Fit Index (CFI \geq 0.90); the Normed Fit Index (NFI \geq 0.90); the Tucker-Lewis Index (TLI \geq 0.90); and the Root Mean Square Error of Approximation (RMSEA \leq 0.08). SEM results were used to compare a single-factor model (which tested all the items of resilience and self-efficacy loading in one factor) with a two-factor model (which tested all the items loading in two different factors) to determine which model fit better. Confirmatory factor analyses (CFA) conducted on the single-factor model revealed inadequate fit indices, χ^2 (44) = 248.59, CFI = 0.76, NFI = 0.73, TLI = 0.64, and RMSEA = 0.16, whereas the two-factor model fit the data well, χ^2 (43) = 90.03, CFI = 0.95, NFI = 0.90, TLI = 0.92, and RMSEA = 0.08. The comparison of χ^2 statistics indicated that the two-factor model fit the data significantly better than the one-factor model $\Delta\chi^2$ (a) = 158.56, p < 0.001. All items had significant loadings in their intended latent factors. Moreover, as expected, self-efficacy was positively and significantly related to AR: β = 0.37, p < 0.001. Based on these findings, we can empirically confirm the factor structure of the resilience measure, and that self-efficacy is significantly related to AR, but they are conceptualized in different ways, thus providing empirical support for the construct validity of the AR-S.

2.2.2 | Reliability

An important step in establishing the psychometric characteristics of a scale is to determine its reliability. We used Cronbach's α as a measure of the internal consistency reliability of the scale because it reflects the extent to which the items are homogeneous. The AR-S showed a Cronbach's α = 0.90 (Kline, 2011), which is excellent.

2.2.3 | Convergent validity

Convergent validity of the scale was checked in two ways. First, we inspected whether our measures of AR were related to another (criterion) resilience measure. The convergent validity of the AR-S was examined by correlating it with the resilience scale from the PCQ-12, adapted to the academic context (Luthans et al., 2008; Luthans, Youssef, Sweetman, & Harms, 2013). Convergent validity is assumed when the scales of the tested instrument correlate positively and moderately to highly with the criterion instrument (Kline, 2011). Pearson's correlation coefficient between the AR-S and the resilience scale from the PCQ-12 showed that the correlation was high and significant (r = 0.58, p < 0.01).

We then computed the average variance extracted (AVE; Fornell & Larcker, 1981). The AVE represents the extent to which items of a specific construct "converge" or share a high proportion of variance in common. A model can be considered to have good convergent validity if at least 50% of the measurement variance is captured by the construct (AVE > 0.50). The estimated AVE of the AR-S is 0.61. Taken together, the results support the convergent validity of the AR-S.

2.2.4 | Divergent validity

Finally, the divergent validity was checked by demonstrating that AR is different from another (criterion) variable. We investigated the divergent validity of the AR-S by correlating it with the proving dimension of goal orientation. A proving goal orientation is a focus on demonstrating one's competence and gaining favorable judgments or positive evaluations from others (Vandewalle, Cron, & Slocum, 2001). Previous evidence showed that proving goal-oriented students typically do not increase their efforts following failures, and that when confronted with obstacles, they exhibit decreased problem-solving and readily disengage from goals, even if they were performing adequately previously (Snyder et al., 2002). Thus, they are not expected to develop resilience when faced with an adverse situation.

Divergent validity is assumed when the scales of the tested instrument show a low-to-moderate correlation with the criterion instrument (Kline, 2011). Pearson's correlation coefficient between the AR-S and the proving goal orientation scale showed that the correlations were not significant (r = -0.01, p = 0.95). These results provide evidence for good divergent validity of the AR-S.

3 | STUDY 2

In the second study, the hypothesized model was explored through the following hypotheses:

Hypothesis 2. Problem-focused coping will be positively related to AR.

Hypothesis 3. Emotion-focused coping will be negatively related to AR.

Hypothesis 4. Meaning-focused coping will be positively related to AR.

Hypothesis 5. AR will be positively related to academic satisfaction.

Hypothesis 6. Academic satisfaction will partially mediate the relationship between AR and performance.

3.1 | Method

3.1.1 | Sample and procedures

The second study involved 780 university students from a Spanish university. The sample size was adequate to carry out the proposed analyses as it exceeded the minimum of 754 observations needed for a statistical power of 0.80 and eight degrees of freedom (MacCallum, Browne, & Sugawara, 1996), the degrees of freedom of our proposed model.

Participants (59.7% females) were stratified and belonged to the four faculties of the University, namely: Faculty of Humanities and Social Sciences (33.2%), School of Technology and Experimental Sciences (25%), Faculty of Law and Economics (24.5%), and Faculty of Health Sciences (17.3%). Ages ranged from 18 to 61 years (M = 22.7; standard deviation [SD] = 5.68 years). Most of them (93.4%) were studying a bachelor's degree lasting 4 years, whereas the others were studying 5-year degrees from the previous curriculum (6.6%). Regarding the year they were studying, 35.7% were in the first year, 37.2% were in the second, 19.6% were in the third, 6.5% were in the fourth, and 9% were in their fifth year. Moreover, 84.1% were not working at the time.

The data were collected through a paper and pencil questionnaire that includes all the study variables. Students received a brief presentation of the study by the researchers during class time, and they were invited to individually

fill out a questionnaire. On the cover page of the questionnaire, they read and signed an individual informed consent. The data collected by the questionnaire and the data on academic performance (GPA provided by the University) were matched through the National Identity document of each student.

3.1.2 | Measures

The students completed the AR-S described in Study 1 to assess their resilience. Moreover, three other measures were used in this study to test our hypotheses.

Coping strategies

The Spanish version of the Brief COPE inventory was used to assess coping strategies (Perczek, Carver, Price, & Pozo, Kaderman, 2000). It includes 14 two-item subscales: active coping, acceptance, emotional support, instrumental support, positive reframing, planning, self-distraction, denial, behavioral disengagement, venting, self-blame, religion, humor, and substance use. An example item is: "In the presence of difficult situations related to my studies, I've been taking action to try to make the situation better" (active coping subscale). For each of the items, respondents indicated the extent to which they used the strategy to deal with stressful situations on a 4-point Likert scale ranging from 0 (not at all) to 3 (a lot). Cronbach's α ranged between 0.62 and 0.94 (mean = 0.77). The order of the coping strategy items was alternated between different subscales.

Academic satisfaction

Academic satisfaction was measured with a three-item scale (Salanova et al., 2005) that takes three salient aspects into consideration: their professors, the degree they are studying for, and the faculty to which they belong. An example item is: "How satisfied are you with your professors?", and students indicated the extent of their satisfaction on a 5-point Face scale ranging from 1 (*frowning*) to 5 (*smiling*). The academic satisfaction score was obtained by computing the mean of the three items.

Academic success

It was assessed by using objective performance as reflected by the GPA, provided by the University at the end of the examination period after the distribution of the questionnaire, that is, 4–5 months later. In accordance with the Spanish system of qualifications, the GPA ranged from 5 (poor) to 10 (excellent).

3.1.3 | Fit indices

Our data were analyzed by means of SEM techniques using the AMOS 21.0 (Arbuckle, 2005). The covariance matrix was analyzed using the maximum-likelihood estimation method. To reduce the complexity of the models examined, we used manifest variables for all constructs (Jöreskog & Sörbom, 1993). To use scores for our coping manifest variables that encapsulate the factor loadings of their underlying subscales, we calculated their weighted factor scores. Because AR, satisfaction, and GPA are one-dimensional constructs, we used their standardized scores as manifest variables (Xanthopoulou, Bakker, Demerouti, & Schaufeli, 2009). Goodness of fit was determined as in Study 1.

3.2 Results

3.2.1 | Preliminary analyses

Before testing our hypotheses, we conducted an exploratory factor analysis (EFA) to examine how the coping subscales grouped together to form broader coping constructs in university students. In fact, the theory proposes that the 14 coping strategies assessed by these subscales are part of several larger constructs (e.g., problem-focused,

emotion-focused, and meaning-focused); however, researchers have warned against the practice of assuming that certain coping strategies are always grouped in the same way across different contexts (Lazarus & Folkman, 1984). Therefore, we standardized the total score of each lower-order coping subscale and conducted a principal axis factoring (PAF) analysis with varimax rotation on the 14 coping subscales. The advantage of PAF is that it takes into account the degree to which each lower-order factor contributes to the overall factor (i.e., factor loadings) when calculating factor scores (Xanthopoulou et al., 2009). Subscales with a loading of less than 0.40 (self-blame, religion, substance use, and self-distraction) were sequentially deleted, which resulted in 10 subscales loading in 4 factors. Factor loadings of the lower-order coping subscales are presented in Table 1. Factor 1 included the planning and active coping subscales, called problem-solving coping. Factor 2 consisted of the emotional support, instrumental support, and venting subscales, called social coping. Factor 3 consisted of the denial and behavioral disengagement subscales, called avoidant coping. Factor 4 included the positive reframing, humor, and acceptance subscales, referred to as positive-acceptance coping. All 4 factors accounted for 45.03% of the total variance explained. In accordance with the coping literature, Factor 1 (problem-solving coping) is consistent with the problem-focused category, Factors 2 and 3 (social and avoidant coping) are representative of the emotion-focused category, and Factor 4 (positive-acceptance coping) is coherent with the meaning-focused category.

To confirm the structure of our coping factors, we also conducted a series of subscale-level CFA. We compared a model where the 14 original subscales from the COPE load in one general coping factor (M1) to a model where the final 10 subscales from the EFA load in one general coping factor (M2), and a model where the final 10 subscales from the EFA load in the respective underlying covariate factors (M3). Results supported the use of 10 coping subscales rather than the 14 originally proposed ($\Delta\chi^2_{\text{M1-M2}}(42) = 243.80$, p < 0.001), and that they belong to the four factors ($\Delta\chi^2_{\text{M2-M3}}(6) = 553.28$, p < 0.001).

3.3 | Descriptive analyses

Table 2 presents the correlations among the study variables, as well as Cronbach's α on the diagonal. Given that the variables used are standardized factor scores, means and standard deviations are not reported. Correlations between the 10 separate coping subscales will be provided by the first author upon request.

To mitigate problems with common method bias, we implemented two procedural remedies, as suggested in Podsakoff, MacKenzie, and Podsakoff (2012). First, we obtained the measures from different sources, specifically

TABLE 1 Results of principal axis factoring analyses: Standardized factor loadings of lower-order coping strategies on higher-order factors (N = 780)

	Higher-order Factors						
Lower-order coping strategies	Problem-solving	Social	Avoidance	Positive acceptance			
Active coping	0.811	-	-	-			
Planning	0.550	-	-	-			
Emotional support	-	0.873	-	-			
Venting	-	0.606	-	-			
Instrumental support	-	0.548	-	-			
Disengagement	-	-	0.655	-			
Denial	-	-	0.642	-			
Positive reframing	-	-	-	0.561			
Humor	-	-	-	0.550			
Acceptance	-	-	-	0.544			

TABLE 2 Correlations for the study variables (N = 780)

	1	2	3	4	5	6	7
1. Social coping	(0.81)	-	-	-	-	-	-
2. Problem-solving coping	-0.06	(0.68)	-	-	-	-	-
3. Positive acceptance coping	-0.04	-0.11**	(0.69)	-	-	-	-
4. Avoidance coping	-0.08*	0.10**	-0.03	(0.69)	-	-	-
5. Academic resilience	-0.26**	0.16**	0.33**	-0.14**	(0.89)	-	-
6. Academic satisfaction	0.01	0.20**	0.01	-0.08*	0.16**	(0.72)	-
7. Academic performance	0.12**	0.12**	-0.06	-0.10**	0.03	0.14**	-

Note. Cronbach's α reliability estimates are listed in the diagonal in parentheses.

the predictor and mediator measures from students and the criterion measure from an objective source. Second, we differentiated the scale properties shared by the measures of the predictor and mediator variables. Finally, the correlations between the study variables were examined, and no significant relationships were found between any of them. When common method bias is present, all of the relationships should be significant; otherwise, the common method bias is so small that it is meaningless (Spector, 2006).

3.3.1 | Test of hypotheses: path analyses

When the model was tested, social and avoidant coping negatively and significantly predicted resilience (β = -0.25, p < 0.001, and β = -0.17, p < 0.001, respectively), whereas problem-solving and positive-acceptance coping positively and significantly predicted resilience (β = 0.20, p < 0.001, and β = 0.34, p < 0.001, respectively). Resilience positively and significantly predicted satisfaction (β = 0.16, p < 0.001), which in turn positively and significantly predicted GPA (β = 0.14, p < 0.001). The relationship between resilience and GPA was not significant (β = 0.01, p = 0.933), which suggests full mediation by satisfaction. The results of the Sobel test confirmed the fully mediated relationship (z = 2.99, p < 0.01). However, our hypothesized model (M1) showed a poor fit to the data: χ^2 (8) = 60.85, RMSEA = 0.09, NFI = 0.82, IFI = 0.84, and CFI = 0.83.

The so-called modification indices indicated that the model fit could be improved significantly by including direct paths from problem-solving coping to academic satisfaction and GPA, as well as a direct path from social coping to GPA. The results of the SEM analyses showed that the modified model (M2) fit the data reasonably well, $\chi^2(6) = 16.15$, RMSEA = 0.05, NFI = 0.95, IFI = 0.97, and CFI = 0.97. M1 and M2 were compared using the χ^2 difference test, and results supported the modified model ($\Delta\chi^2_{\text{M1-M2}}(3) = 44.7$, p < 0.001). The relationships between the variables were almost invariant between M1 and M2, but the inclusion of these three paths significantly improved the model. Moreover, these additional relationships in M2 were significant, showing that problem-solving coping has a direct impact on academic satisfaction and that both problem-solving and social coping have a direct impact on GPA. Figure 1 shows the final model with the standardized coefficients.

4 | DISCUSSION

The lack of specific tools to measure resilience in academic contexts in Spain makes it important to validate suitable instruments for studying AR. These instruments allow resilience to be evaluated and related to other variables. In this regard, the purpose of the present study was twofold: (a) to evaluate the psychometric characteristics of the AR-S by Martin and Marsh (2006) in the Spanish context, checking for its validity and reliability; and (b) to

^{*}p < 0.05.

^{**}p < 0.01.

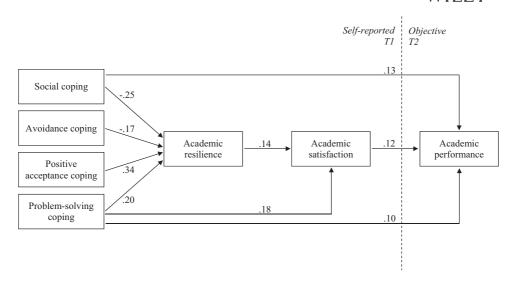


FIGURE 1 The final model with standardized path coefficients (N = 780)

investigate the relationships between coping strategies, AR, and academic satisfaction, as well as the impact on academic performance over time. To achieve these objectives, we conducted two studies. The findings of the first study suggest that the Spanish version of the AR-S demonstrated acceptable psychometric properties. Regarding the second study, the results revealed that students who used emotion-focused coping strategies showed less AR, whereas resilience was higher when problem- and meaning-focused coping strategies were used. Moreover, students with a high level of AR tended to be more satisfied with their academic context, and students who displayed more satisfaction tended to perform better over time. Furthermore, results indicated that AR was only indirectly related to performance, through satisfaction. Finally, the study revealed three unpredicted significant paths linking coping strategies to outcomes, namely, problem-solving coping to both academic satisfaction and performance, and social coping to academic performance.

4.1 | Theoretical contributions

The present study represents a step forward with respect to previous research on AR in several ways. First, the validation of the Spanish version of the AR-S by Martin and Marsh (2006) was successful. To provide solid evidence for the psychometric characteristics of the scale, we carried out a number of analyses, rigorously following well-established validation procedures, and we can conclude that the Spanish version of the AR-S demonstrates good psychometric properties. A contribution of the study is the use of a heterogeneous sample, and so the Spanish translation of the AR-S can be used by scholars as a valid scale to investigate resilience in the academic context.

Second, we advance the theoretical understanding of the relationship between different kinds of coping strategies and AR. One strength of the current study is that we factor analyzed the subscales to allow student responses to determine subscale categorization. However, the factors found in this study are strongly supported by previous theoretical conceptualizations (Folkman & Moskowitz, 2004), thus providing support for the validity of our categories. Specifically, we have identified two coping strategies, problem-solving and positive acceptance, that are positively related to resilience. Within the academic context, problem-solving coping should assist students in adjusting to the many challenges posed by the environment as it involves taking active steps to alter the circumstances or address the problem. Therefore, students who tend to change the situation when they are faced with stressful events have higher levels of resilience. Positive-acceptance coping underlies the cognitive strategies used to manage the meaning of a situation and regulate positive emotions, which is consistent with the emerging

focus on meaning-focused coping (Folkman, 2008). In turn, meaning-focused coping is said to generate positive emotions and the underlying appraisals, and these emotions and appraisals influence the resilience process. Therefore, students who tend to generate positive emotions by attributing meaning to the situation show higher resilience. By contrast, we have identified two coping strategies, social and avoidance coping, that are negatively related to resilience. Both strategies belong to the traditional emotion-focused category and have to do with the regulation of negative emotions, although each of them refers to a definite type of strategy. Through social coping, students look for social support and vent emotions. Although the strategy can help to reduce stress at first, symptom reduction can be dysfunctional if students mentally focus too much on the situation that created the stress when venting. Therefore, students who use more social coping strategies show lower resilience. Finally, avoidance is a coping strategy students that use to try to ignore a stressful situation by not thinking about it. Ignoring a stressful situation can be positive as it might reduce stress, but avoiding the reality of a stressful situation can allow the situation to worsen, thus increasing stress in the long run (Belanger, Lewis, Kasper, Smith, & Harrington, 2007). Hence, students who tend to avoid the problem have lower levels of resilience.

Third, in studying the positive outcomes of AR, we further extend current theory in at least two ways. On the one hand, past research on resilience outcomes focused predominantly on successful performance as an indicator of positive adjustment. However, students' attitudes and beliefs are more proximal outcomes (Baker et al., 2003) and may be better able to predict eventual academic-related behaviors, such as performance. Thus, in this study, we included academic satisfaction as a measure of positive academic outcomes, and we found that AR is positively related to satisfaction. The role of AR in predicting satisfaction seems particularly encouraging because it suggests that a higher level of resilience is associated with increased satisfaction with their academic context. On the other hand, in contrast to our expectations, we found that resilience is only associated with performance via satisfaction. Although further investigation is needed, this result seems to refute the recognized statement that higher resilience is a predictor of better performance (Luthar, 1991). Some possible explanations can be provided for this unexpected result. Although resilience may influence academic performance, so do many other circumstances and student characteristics, such as performance self-efficacy, grade goal, and effort regulation (Ouweneel, Schaufeli, & Le Blanc, 2013; Richardson, Abraham, & Bond, 2012). Moreover, AR might not have a direct impact on performance because the process of bouncing back is not necessarily directly related to an increase in performance. For instance, a psychological process engendered by resilience—through an increase in attitudes (i.e., satisfaction)—affected performance.

Finally, unexpected results revealed the direct effects of the selected coping strategies on outcomes. Specifically, we found a direct relationship that links problem-solving coping to both academic satisfaction and performance, thereby providing evidence that the cumulative effect of using problem-solving strategies across a range of stressful situations results in better outcomes. Because problem-solving coping involves strategies to alter or diminish a stressful event, it seems plausible that it would be positively related to satisfaction and performance. For instance, if the stressor is an impending assignment, focusing on the assignment should result in getting the assignment handed in on time and getting a good grade (MacCann, Fogarty, Zeidner, & Roberts, 2011). Furthermore, we found a direct relationship between social coping and performance. There are several possible pathways through which social coping strategies might influence academic performance. First, students who are able to regulate their negative emotions through venting or support seeking could be less impaired in assessment and learning situations. Second, in the Spanish context, academic performance requires not only passing examinations, but also collaboration and teamwork, for instance, in the form of group projects (Ahles & Bosworth, 2004). Social coping has been linked to better social relationships, suggesting that individuals who use this kind of coping might be better able to maintain the social relationships required for effective group work and, thus, achieve higher grades (MacCann et al., 2011).

4.2 | Practical implications

In terms of practical implications, we can conclude that the AR-S is a valid and reliable instrument for measuring resilience in the Spanish academic context. Additionally, this questionnaire contains only six items to measure

resilience and, therefore, is short and practical. Thus, the AR-S can be considered a solid tool for conducting research on AR.

Another important point underpinning the present study is related to the proposition that both coping strategies and AR can be influenced in a direct way. Thus, a practical implication of the study is that, in the university context, both coping strategies and resilience can be used as tools to boost students' outcomes, especially for those who are underperformers or have low satisfaction. Given that our results suggest that AR is enhanced by specific coping strategies, interventions aimed at encouraging problem-solving and positive-acceptance coping strategies seem to be crucial. In this regard, there are several interventions designed to teach individuals how to cope with adversity. They typically include techniques that help the individual to deal with and manage stress (see MacCann et al., 2011).

Moreover, evidence shows that resilience can also be developed through interventions. For example, repeated exposure to positive emotions may help to strengthen an individual's resilience capacity (Fredrickson, Tugade, Waugh, & Larkin, 2003). Specifically, positive emotions may take the form of laughter or smiling, and such emotions may reinforce or strengthen resilience (Bonanno, Noll, Putnam, O'Neill, & Trickett, 2003). Although these types of positive emotions may seem simplistic, their effects can be important. Taken together, these guidelines suggest that coping and resilience could be modifiable, leading to enhanced positive academic outcomes.

4.3 | Limitations and research directions

This study has several limitations that highlight important avenues for future research. With the exception of academic performance rates (i.e., GPA), perhaps the clearest limitation is the use of self-reported data, which increases the risk of common method bias (Podsakoff, MacKenzie, Lee, & Podsakoff, 2003). Although the use of self-reports in this study could be justified by the nature of the constructs because students are the most accurate source of their own strategies and feelings or attitudes, it introduces the possibility of response acquiescence and precludes discussion of causality. We tried to minimize these errors by differentiating the response scales for each of these variables, as suggested in Podsakoff et al. (2012).

Another limitation of the present study is that the data are cross-sectional, with the exception of the GPA, which is collected at a different point in time. Although SEM analysis provides some information about the possible direction of the relationships, cross-sectional study designs do not allow us to draw firm conclusions about the causal ordering among the variables studied. Clearly, there is a need for longitudinal studies to strengthen causal inferences about the relationships among coping strategies, resilience, and a range of academic outcomes.

In addition, although the sample came from four different faculties and several degrees, our results are based on a sample from the same university. Thus, the results need to be replicated to generalize our findings to different academic contexts.

4.4 | Conclusions

Although research on AR is usually focused on students who encounter greater adversity, it is important to recognize its importance for university students in general because they all have to deal with adverse and stressful situations during their academic experience. The results of the current study indicate that the Spanish translation of the AR-S can be used in a valid and reliable way to measure AR. Thus, this measure can be used to investigate the relationship between AR and relevant antecedents and outcomes. In this regard, our results offer new insights into the way university students can benefit from using adequate coping strategies to support the development of their resilience and achieve positive outcomes, such as satisfaction and performance. These findings have implications for researchers investigating the processes students use to deal with everyday setbacks, and for practitioners seeking to assist students in dealing with the highs and lows that characterize academic life.

ACKNOWLEDGMENTS

This study was supported by the Fundació Caixa Castelló-Bancaixa under Grant E 2013-31 and Consellería de Educación Generalitat Valenciana under Grant PROMETEO/2013/025.

CONFLICT OF INTERESTS

The authors declare that there are no conflict of interests.

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How to cite this article: Meneghel I, Martínez IM, Salanova M, de Witte H. Promoting academic satisfaction and performance: Building academic resilience through coping strategies. *Psychol Schs.* 2019;56:875–890. https://doi.org/10.1002/pits.22253